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## Lorenzo Frattarolo, Francesca Parpinel, and Claudio Pizzi

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# Lorenzo Frattarolo Francesca Parpinel Claudio Pizzi 

Ca' Foscari University of Venice, Department of Economics


#### Abstract

According to the definition of Financial Stability Board (FSB), Systemically Important Banks (SIBs) are the banks "whose disorderly failure, because of their size, complexity and systemic interconnectedness, would cause significant disruption to the wider financial system and economic activity". The current methodology for their determination is based on balance-sheet variables and expert judgment. We use permutation tests to investigate the relevance of equity-based systemic risk measures in the SIBs choice. Restriction of the analysis to European Banks, for which full information is available, allows understanding the importance of equity-based systemic risk measures also for size, interconnectedness, substitutability/financial Institution Infrastructure, complexity and cross-jurisdictional Activity categories.


Keywords: Systemic risk, Particle Swarm Optimization, Nonparametric Combination
JEL Codes: C14, C61, C63, G21, G28

> Address for correspondence:
> Lorenzo Frattarolo
> Department of Economics
> Ca' Foscari University of Venice
> Cannaregio 873, Fondamenta S.Giobbe
> 30121 Venezia - Italy
> Fax: (++39) 0412349176
> E-mail: lorenzo.frattarolo@unive.it

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## 1. Introduction and Objectives

The framework of Systemically Important Financial Institutions (SIFIs) was introduced by the Financial Stability Board (FSB) in October of 2010 as the institutions "whose disorderly failure, because of their size, complexity and systemic interconnectedness would cause significant disruption to the wider financial system and economic activity", FSB (2010). The current methodology to determine the Globally Systemically Important Banks (G-SIBs) is outlined by the Bank for International Settlements, BIS (2013). In particular, the banks included in the analysis have to fulfill any of the following criteria:

- Banks that the Committee identifies as the 75 largest global banks, according to the leverage ratio exposure measure, based on Basel III and at the end of the financial year.
- Banks that were designated as G-SIBs in the previous year (unless supervisors agree that there is a compelling reason to exclude them).
- Banks, with a score produced by the indicator-based measurement approach that exceeds a cut-off level set by the Committee.
- Banks that have been added to the sample by national supervisors using supervisory judgment (subject to certain criteria).

The regulator builds the selection process on proprietary annual data. The collection of the dataset is time-consuming and delays the publication of the selection. For example, the FSB published the last release based on end-2014 data only in November 2015. The European Banking Union provides full disclosure of the data used to define the European SIBs for the year 2014 (using data from 2013) and 2015 (using data from 2014). For a review of the literature on the G-SIBs and a critique of the methodology see Iwanicz-Drozdowska (2014) and Barth, Nolle, Li, and Brummer (2014). Bongini, Nieri, and Pelagatti (2015) discuss the financial impact of the SIFIs selection. In parallel, several papers proposed systemic risk measure based on stock returns. For a clear description of those measures refer to the recent survey Bisias, Flood, Lo, and Valavanis (2012). In this way, market participants are able to compute them at any moment due to their easy access to this type of data. Furthermore, Scholars use them in ranking timely the systemic importance of single institutions. Recently, Giglio, Kelly, and Pruitt (2016) investigate their macroeconomic meaning and substantiated this practice. The study of their relationship with the variables used by the FSB is new in the literature to our knowledge. A bridge between those two set of measures could bring a more timely and accurate choice of the SIBs. All the above reasons justify the present investigation of the linkage among them. In particular, we consider as given the choice made by the FSB about the two groups of SIBs and non-SIBs. Combining statistics from permutations of several variables is a well-known possibility, see Pesarin and Salmaso (2010). We propose a combination based on the systemic risk measures to test the identity of the two FSB groups. Rejection of the null hypothesis implies the reproduction of the choice using equity information. So, we choose the combination weights by minimizing the $p$-value of the combination test. If the weights add up to one, one can interpret them as the measure relevance in the FSB choice. Also, we are able to associate a SIBs selection to each macro category using a heuristic. Then, we repeat the procedure for size, interconnectedness, substitutability, complexity and cross-jurisdictional Activity. Non-smoothness of the $p$-value as a function of weights requires a global optimization method. The parallelizable Particle Swarm Optimization (PSO) is valid candidate, see Kennedy (2010). In fact, we exploit the multi-core architecture of the Sistema
per Calcolo Scientifico di Ca' Foscari (SCSCF) for computations ${ }^{1}$. The paper is organized as follows. In Section 2 the methodology is presented, in Section 3 we show the results for the FSB selection and for macro categories driven selections, then in Section 4 we discuss the results and propose some possible extensions.

## 2. Methodology

In this Section, we review the permutation and combination test methodology, following Pesarin and Salmaso (2010), introduce our optimization procedure and justify the use of Particle Swarm technique. Let be $I(\cdot)$ the indicator function, that is equal to 1 if the condition in parenthesis is satisfied and zero otherwise.
The baseline permutation test for a single cross sectional variable, $X$, can be summarized in the following steps:

1. The observed units, corresponding to different banks, are divided into two groups, $g_{1}$ and $g_{2}$, according to a given selection.
2. Given a cross sectional variable, $X$, observed on values $x_{i}, i=1, \ldots, n$, compute a relevant statistic for each group; in our case we use the empirical distribution function $\hat{F}_{k}$ for group $k=1,2$. The maximum of the difference of the two functions $K S\left(\hat{F}_{1}, \hat{F}_{2}\right)$, that is a two-sample Kolmogorov-Smirnov statistic, will be our test statistic, and its observed value will be $v_{o b s}$.
3. Exchange randomly the participants in the groups, retaining only their sizes. We randomly choose a permutation of the indices $i$, named $\pi_{b}, b=1, \ldots, B$, obtaining two new groups $g_{1}^{b}$ and $g_{2}^{b}$. Then, considering the exchangeability assumption of $X$ and under the hypothesis of identical distribution for the two groups, $H_{0}: F_{1}=F_{2}$, the statistic $v^{b}=K S\left(\hat{F}_{g_{1}^{b}}, \hat{F}_{g_{2}^{b}}\right)$ would have the same distribution of $v_{\text {obs }}$.
4. Compute, according to Pesarin and Salmaso (2010), an approximated $p$-value by

$$
\begin{equation*}
P_{B}=\frac{1}{B} \sum_{b=1}^{B} 1\left(v_{b} \geq v_{o b s}\right) \tag{1}
\end{equation*}
$$

Figure 1 shows an example on permutation test for a simulation of $n=500$ observations from a bimodal distribution.
The combination of several partial permutation tests requires additional steps. After the definition of the size of each group, we apply the procedure outlined before to get the permutation distribution of each partial test statistic $t_{j}$, typically $t_{j, \text { obs }}=K S\left(\hat{F}_{1}^{j}, \hat{F}_{2}^{j}\right), j=1, \ldots, p$, where the empirical cumulative distribution functions, $\hat{F}_{g_{k}}^{(j)}, k=1,2$, refer to the observed values of $X_{j}$, furthermore we denote with $t_{j}^{b}$ the partial test statistics computed on each permutation $b$ of the two groups, with $b=1, \ldots, B$. Then each dimension is transformed to an auxiliary variable related into the single $p$-values

$$
\begin{equation*}
\lambda_{j}=\frac{1}{B+1}\left(\frac{1}{2}+\sum_{b=1}^{B} 1\left(t_{j}^{b} \geq t_{j, o b s}\right)\right) \tag{2}
\end{equation*}
$$

[^0]

Figure 1: Explanation of permutation test
that takes values strictly inside the unit interval. So, it may be defined in such a way that they can be aggregated in a single variable using a combination function. In our case, we use a Fisher omnibus function with the same weights of the indices $t=-\sum_{j=1}^{p} w_{j} \log \left(\lambda_{j}\right)$. In addition, we get the value of the statistics in each permutation, by

$$
\begin{equation*}
\lambda_{j}^{b}=\frac{1}{B+1}\left(\frac{1}{2}+\sum_{b=1}^{B} 1\left(t_{j}^{r} \geq t_{j}^{b}\right)\right) \tag{3}
\end{equation*}
$$

in such a way to produce an approximated permutation distribution. This procedure may be extended to the combined variable for each permutation, $t^{b}=-\sum_{j=1}^{p} w_{j} \log \left(\lambda_{j}^{b}\right)$. Given $B$ random permutations, according to Pesarin and Salmaso (2010), we can obtain an approximated $p$-value by

$$
\begin{equation*}
P^{B}=\frac{1}{B} \sum_{b=1}^{B} 1\left(t_{b} \geq t\right) \tag{4}
\end{equation*}
$$

and we can reject the global null hypothesis of equality the two groups $H_{0}: F_{1}=F_{2}$ at $\alpha$ significance level if $P^{B} \leq \alpha$.
The lower is the value of $P^{B}$ the higher is the significance of the test and the higher is the difference between the multivariate distribution of the two groups. For this reason we choose to optimize the weights in order to minimize the $p$-value.

### 2.1. Particle Swarm Optimization

In order to test our procedure we consider an optimization problem in which the function that have to be optimized depends on the weights used in the Fisher omnibus function that combines several tests. Given the complexity of the optimization problem we consider the bio-inspired iterative metaheuristic called Particle Swarm Optimization (PSO) introduced by Kennedy and Eberhart (1995). This procedure, based on swarm intelligence, is a robust stochastic method for unconstrained optimization problem although it is possible to treat also constrained ones, seeCorazza, Fasano, and Gusso (2013). The PSO algorithm exploits the concept of social intelligence and co-operation to mimic the intelligence that moves together individuals of the same species looking for food.
To this aim, each member of the swarm explores the search area reminding its best position reached so far, exchanging the information with the other ones. The whole swarm will converge to the best global position.
Each member of the swarm is a particle and represents a possible solution to the investigated optimization problem. The initialization of the particle placement is random as well as their initial velocity. The movement of each particle of the swarm depends on its position with respect to the best position. In the following, we give a description of the standard PSO procedure.
Let us consider the unconstrained optimization problem

$$
\min _{x \in R^{d}} f(x)
$$

where $f: R^{d} \rightarrow R$ is the objective function in the minimization problem. At the $k$-th iteration of the PSO algorithm, the following objects are associated to the $j$-th particle $\{j=1, \ldots, M\}$ :

- $x_{j}^{k} \in R^{d}$ : position of the $j$-the particle at step $k$;
- $f\left(x_{j}^{k}\right) \in R$ : fitness of the $j$-th particle at step $k$;
- $v_{j}^{k} \in R^{d}$ : velocity of the $j$-the particle at step $k$;
- $p_{j} \in R^{d}$ : best position visited so far by the $j$-th particle at step $k$.

Furthermore pbest $_{j}$ denotes the objective function in the position $p_{j}$ of the $j$-th particle.

- Set $k=0$. Randomly generate $x_{j}^{k}$ and $v_{j}^{k}$.
- Set $k=1$. Set pbest ${ }_{j}=+\infty$ for all $j$ and set gbest $=+\infty$.

The overall PSO algorithm is described in the following.

1. Evaluate $f\left(x_{j}^{k}\right)$ for all $j$.
2. If $f\left(x_{j}^{k}\right)<$ pbest $_{j}$ then $p_{j}=x_{j}^{k}$ and pbest ${ }_{j}=f\left(x_{j}^{k}\right)$ for all $j$. If $f\left(x_{j}^{k}\right)<g b e s t$ then $p_{g}=x_{j}^{k}$ and gbest $=f\left(x_{j}^{k}\right)$.
3. Update velocity and position for all $j$ :

$$
\begin{gathered}
v_{j}^{k+1}=w^{k+1} v_{j}^{k}+U_{1} \otimes\left(p_{j}-x_{j}^{k}\right)+U_{2} \otimes\left(p_{g}-x_{j}^{k}\right) \\
x_{j}^{k+1}=x_{j}^{k}+v_{j}^{k+1}
\end{gathered}
$$

where $U_{1}, U_{2} \in R^{d}$ such that each their dimension is a realization of uniform random variables over $\left[0, \theta_{1}\right]$ and $\left[0, \theta_{2}\right]$, respectively. The symbols $\otimes$ denotes the componentwise product and $p_{g}$ is the best position of the $j$-th particle in a neighborhood. The value assumed by the inertia weight $w^{k}$ will be discussed in the following.
4. If a convergence criterion is not satisfied then update $k=k+1$ and go to step 2 .

The convergence criterion is either the maximum number of iterations or if there are no movements among particle.
The choice of the inertia weights $w^{k}$ affects the convergence of the swarm. In literature, the parameters $w^{k}, k=1, \ldots, K$ are generally proposed as linearly decreasing together with the number of steps, i.e.

$$
w^{k}=w_{\max }+\frac{w_{\max }-w_{\min }}{K} \cdot k
$$

where common values of $w_{\max }$ and $w_{\min }$ are respectively 0.9 and 0.4 , while $K$ is usually the maximum number of allowed steps.
One of the great advantages of the particle swarm optimization is that it is highly parallelizable. In fact, the updating of particles characteristics is independently computable given the previous iteration. As already remarked, the SCSCF cluster performed the computations. This allows us to use 200 hundred particles handled by 100 cores of the cluster. Instead one of the main drawbacks of particle swarm optimization is that is stochastic in nature. In our case also the permutation test is randomized doubling the source of stochasticity. Robustness

Table 1: Systemic Measures and Indicators: correlations

|  | Kendall's $\tau 2013$ |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | -MES | - $\Delta$ CoVaR | in connections | out connections | in out connections | 1/closeness | eigenvector centrality | PCA |
| Size | $\begin{gathered} -0.01 \\ (0.9816) \end{gathered}$ | $\begin{gathered} 0.20 \\ (0.1699) \end{gathered}$ | $\begin{gathered} -0.15 \\ (0.3693) \end{gathered}$ | $\begin{gathered} -0.04 \\ (0.8419) \end{gathered}$ | $\begin{gathered} -0.13 \\ (0.4229) \end{gathered}$ | $\begin{gathered} -0.01 \\ (0.9808) \end{gathered}$ | $\begin{gathered} -0.05 \\ (0.7618) \end{gathered}$ | $\begin{gathered} -0.04 \\ (0.7993) \end{gathered}$ |
| Interconnectedness | $\begin{gathered} -0.03 \\ (0.8715) \end{gathered}$ | $\begin{gathered} 0.23 \\ (0.1076) \end{gathered}$ | $\begin{gathered} -0.11 \\ (0.5167) \end{gathered}$ | $\begin{gathered} -0.06 \\ (0.727) \end{gathered}$ | $\begin{gathered} -0.12 \\ (0.4515) \end{gathered}$ | $\begin{gathered} -0.07 \\ (0.648) \end{gathered}$ | $\begin{gathered} -0.07 \\ (0.6861) \end{gathered}$ | $\begin{gathered} -0.07 \\ (0.6273) \end{gathered}$ |
| Substitutability | $\begin{gathered} 0.00 \\ (1.000) \end{gathered}$ | $\begin{gathered} 0.09 \\ (0.5631) \end{gathered}$ | $\begin{gathered} 0.08 \\ (0.6535) \end{gathered}$ | $\begin{gathered} 0.02 \\ (0.9205) \end{gathered}$ | $\begin{gathered} 0.06 \\ (0.7156) \end{gathered}$ | $\begin{gathered} -0.07 \\ (0.648) \end{gathered}$ | $\begin{gathered} 0.04 \\ (0.8006) \end{gathered}$ | $\begin{gathered} -0.09 \\ (0.5631) \end{gathered}$ |
| Complexity | $\begin{gathered} -0.06 \\ (0.6943) \end{gathered}$ | $\begin{gathered} 0.15 \\ (0.3187) \end{gathered}$ | $\begin{gathered} -0.04 \\ (0.8419) \end{gathered}$ | $\begin{gathered} -0.03 \\ (0.881) \end{gathered}$ | $\begin{gathered} -0.04 \\ (0.7893) \end{gathered}$ | $\begin{gathered} 0.00 \\ (1.000) \end{gathered}$ | $\begin{gathered} -0.10 \\ (0.5444) \end{gathered}$ | $\begin{gathered} -0.09 \\ (0.5322) \end{gathered}$ |
| Cross-Jurisdictional Activity | $\begin{gathered} 0.06 \\ (0.6943) \end{gathered}$ | $\begin{gathered} 0.28 \\ (0.0519) \end{gathered}$ | $\begin{gathered} -0.20 \\ (0.2124) \end{gathered}$ | $\begin{gathered} -0.09 \\ (0.5832) \end{gathered}$ | $\begin{gathered} -0.23 \\ (0.1385) \end{gathered}$ | $\begin{gathered} -0.08 \\ (0.6138) \end{gathered}$ | $\begin{gathered} -0.06 \\ (0.7236) \end{gathered}$ | $\begin{gathered} -0.09 \\ (0.5322) \end{gathered}$ |
|  | -MES | - $\Delta$ CoVaR | in connections | Kendal out connections | s $\tau 2014$ <br> in out connections | 1/closeness | eigenvector centrality | PCA |
| Size | $\begin{gathered} 0.03 \\ (0.8715) \end{gathered}$ | $\begin{gathered} 0.16 \\ (0.2755) \end{gathered}$ | $\begin{gathered} 0.22 \\ (0.1582) \end{gathered}$ | $\begin{gathered} 0.04 \\ (0.7907) \end{gathered}$ | $\begin{gathered} 0.07 \\ (0.637) \end{gathered}$ | $\begin{gathered} 0.09 \\ (0.583) \end{gathered}$ | $\begin{gathered} -0.07 \\ (0.7066) \end{gathered}$ | $\begin{gathered} -0.05 \\ (0.7288) \end{gathered}$ |
| Interconnectedness | $\begin{gathered} -0.01 \\ (0.9447) \end{gathered}$ | $\begin{gathered} 0.12 \\ (0.4176) \end{gathered}$ | $\begin{gathered} 0.32 \\ (0.041) \end{gathered}$ | $\begin{gathered} -0.01 \\ (0.9808) \end{gathered}$ | $\begin{gathered} 0.09 \\ (0.5396) \end{gathered}$ | $\begin{gathered} -0.07 \\ (0.6501) \end{gathered}$ | $\begin{gathered} -0.13 \\ (0.4516) \end{gathered}$ | $\begin{gathered} -0.07 \\ (0.6605) \end{gathered}$ |
| Substitutability | $\begin{gathered} -0.07 \\ (0.6273) \end{gathered}$ | $\begin{gathered} 0.07 \\ (0.6273) \end{gathered}$ | $\begin{gathered} 0.19 \\ (0.2238) \end{gathered}$ | $\begin{gathered} -0.11 \\ (0.4841) \end{gathered}$ | $\begin{gathered} -0.07 \\ (0.637) \end{gathered}$ | $\begin{gathered} -0.01 \\ (0.9429) \end{gathered}$ | $\begin{gathered} -0.23 \\ (0.1675) \end{gathered}$ | $\begin{gathered} -0.13 \\ (0.3914) \end{gathered}$ |
| Complexity | $\begin{gathered} -0.10 \\ (0.5022) \end{gathered}$ | $\begin{gathered} 0.09 \\ (0.5631) \end{gathered}$ | $\begin{gathered} 0.20 \\ (0.2059) \end{gathered}$ | $\begin{gathered} -0.15 \\ (0.3226) \end{gathered}$ | $\begin{gathered} -0.02 \\ (0.8874) \end{gathered}$ | $\begin{gathered} -0.08 \\ (0.6161) \end{gathered}$ | $\begin{gathered} -0.24 \\ (0.149) \end{gathered}$ | $\begin{gathered} -0.17 \\ (0.2554) \end{gathered}$ |
| Cross-Jurisdictional Activity | $\begin{gathered} -0.01 \\ (0.9447) \end{gathered}$ | $\begin{gathered} 0.17 \\ (0.2363) \end{gathered}$ | $\begin{gathered} 0.31 \\ (0.0461) \end{gathered}$ | $\begin{gathered} 0.05 \\ (0.7538) \end{gathered}$ | $\begin{gathered} 0.11 \\ (0.4502) \end{gathered}$ | $\begin{gathered} 0.01 \\ (0.9429) \end{gathered}$ | $\begin{gathered} -0.08 \\ (0.6606) \end{gathered}$ | $\begin{gathered} -0.17 \\ (0.2363) \end{gathered}$ |

checks reported in the next Section ensure reproducibility of the whole procedure. The use of 200 particles and 10000 random permutations suffices to get reliable results.

## 3. Empirical results

In this section, we begin introducing the used data. Then we derive the selection associated with the macro categories. Finally, we report and comment our results on the optimal weights obtained for 2013 and 2014.
As a preliminar result we can see the correlations among the indicators used by the Regulator and the Systemic measures suggested in literature: Table 1 report the Kendall $\tau$ correlations between the pairs of variables denoting that only few relations are significative, i.e. Cross-Jurisdictional Activity with $\Delta$ Covar in 2013 and Cross-Jurisdictional Activity and Interconnectedness with in connections in 2014.

### 3.1. G-SIB score evaluation and data

Regulators chose SIBs using a composite indicator that considers different aspects of systemic risk: Size, Interconnectedness, Substitutability, Complexity and Cross-Jurisdictional Activity. Those macro categories are themselves composite. The finest subdivision comprises a total of 12 variables detailed in Table 2 BIS (2014). Then a weighted sum of the indicators with the weights as in the last column of Table 1 represents the score.

Table 2: Indicators and relative score weights used by the Basel Committee for the evaluation of the Systemically Important Banks

| Category | Indicator | Indicator weight |
| :--- | :--- | :---: |
| Size | Total exposures | $1 / 5=20 \%$ |
| Interconnectedness | Intra-financial system assets | $1 / 15=6.66 \%$ |
|  | Intra-financial system liabilities | $1 / 15=6.66 \%$ |
| Substitutability/financial institution | Securities outstanding | $1 / 15=6.66 \%$ |
| institution infrastructure | Payment activity | $1 / 15=6.66 \%$ |
|  | Assets under custody | $1 / 15=6.66 \%$ |
|  | Underwritten transactions in debt and |  |
| Complexity | equity markets | $1 / 15=6.66 \%$ |
|  | National amount of OTC derivatives | $1 / 15=6.66 \%$ |
|  | Trading and AFS securities | $1 / 15=6.66 \%$ |
| Cross-jurisdictional | Level 3 assets | $1 / 15=6.66 \%$ |
| activity | Cross-jurisdictional claims | $1 / 10=10 \%$ |
|  | Cross-jurisdictional liabilities | $1 / 10=10 \%$ |

The BIS reports balance sheets included in the SIBs 2014 evaluation sample. The extraction of the relevant indicators from them would be a nontrivial task. So we focus on the European sample whose variables are available on the EBA website. Here EBA publishes balance sheets variables in manageable excel format for 2013 and 2014. In 2013 within a pool of 36 European Banks, 13 of them were considered SIBs. In 2014 within a pool of 37 , 14 of them were chosen as SIBs. To get alternative selection based on the 5 macro categories we proceed in the following way. We consider values of the category corresponding to the institutions selected by the regulator. We compute the minimum value among them. We include all the institutions with values higher than the minimum in the new selection. We report in Table 3 the sizes in the regulator sample and the percentage of institutions still in the selection. As you can see selection sets obtained in this way always contain the regulator set. Unfortunately, not all the considered European banks are listed companies. Stock prices are available for 25 of them. We are able to compute systemic risk measures only for them. In Table 7 and 8 in Appendix we report the names of the used institutions and the computation of the considered systemic risk measure. For a description of the Marginal Expected Shortfall (MES) we refer to Acharya, Pedersen, Philippon, and Richardson (2010). We did not use their SES measure because it includes also information about leverage. The $\Delta \mathrm{CoVaR}$ is introduced and described in Adrian and Brunnermeier (2016). The remaining Interconnectedness measures come from Billio, Getmanski, Lo, and Pelizzon (2012). For a description of several other measures using also a different type of data refer to Bisias et al. (2012). We consider the opposite of MES and $\Delta \mathrm{CoVaR}$ and the inverse of the closeness to obtain comonotonicity among the measures. With this convention, an higher measure corresponds to higher risk.

### 3.2. Systemic Measures contribution to Selections

Table 3 summarizes the main findings of this work. For each SIB selection rows report the optimal $p$-value and the optimal weights in 2013 and 2014. In Section 2 we gave a plausible interpretation of the optimal weights. They represent the share of information about the selection associated with each measure. Lower $p$-values lead to a higher probability that measures with high weights are informative about the selection. In this regards, in

Table 3: Systemic Measures contribution to Selections

| 2013 |  |  |  | weights |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | $\begin{array}{\|l\|} \hline \text { SIBs } \\ \text { partial } \end{array}$ | $\begin{gathered} \text { SIBs } \\ \text { included } \end{gathered}$ | pval | MES | $\Delta \mathrm{CoVaR}$ | $\begin{array}{r} \text { in } \\ \text { degree } \end{array}$ | $\begin{array}{r} \text { out } \\ \text { degree } \end{array}$ | in out degree | closene | eigenvector centrality | PCA |
| Regulator Selection |  | $\mathrm{n}=13$ | 0.002 | 0.001 | 0.572 | 0.023 | 0.001 | 0.378 | 0.023 | 0.001 | 0.001 |
| Size | 23 | 100\% | 0.1841 | 0.013 | 0.0000 | 0.0000 | 0.0000 | 0.9862 | 0.0004 | 0.0001 | 0.0004 |
| Interconnectedness | 19 | 100\% | 0.012 | 0.1320 | 0.3755 | 0.0000 | 0.0001 | 0.3450 | 0.0000 | 0.0000 | 0.1473 |
| Substitutability/Financial | 20 | 100\% | 0.016 | 0.0000 | 0.7884 | 0.0016 | 0.0564 | 0.0050 | 0.1486 | 0.0000 | 0.0000 |
| Complexity | 23 | 100\% | 0.053 | 0.0000 | 0.2010 | 0.0000 | 0.0000 | 0.0000 | 0.0000 | 0.0117 | 0.7873 |
| CrossJurisdictional Activity | 23 | 100\% | 0.003 | 0.0016 | 0.3791 | 0.0973 | 0.0922 | 0.0000 | 0.0000 | 0.0247 | 0.4051 |
| 2014 |  |  |  | weights |  |  |  |  |  |  |  |
|  | $\begin{aligned} & \mathrm{SIBs} \\ & \text { partial } \end{aligned}$ | $\begin{gathered} \text { SIBs } \\ \text { included } \end{gathered}$ | pval | MES | $\Delta \mathrm{CoVaR}$ | degree | $\begin{array}{r} \text { out } \\ \text { degree } \end{array}$ | in out degree | closene | eigenvector centrality | PCA |
| Regulator Selection |  | $\mathrm{n}=12$ | 0.312 | 0.0000 | 0.0000 | 0.0000 | 0.0000 | 0.0211 | 0.0001 | 0.9782 | 0.0006 |
| Size | 16 | 100\% | 0.091 | 0.0005 | 0.5900 | 0.3546 | 0.0004 | 0.0000 | 0.0000 | 0.0000 | 0.0545 |
| Interconnectedness | 14 | 100\% | 0.190 | 0.0000 | 0.0000 | 0.9577 | 0.0000 | 0.0000 | 0.0422 | 0.0001 | 0.0000 |
| Substitutability/Financial Institution Infrastructure | 16 | 100\% | 0.015 | 0.0000 | 0.0002 | 0.9471 | 0.0024 | 0.0000 | 0.0018 | 0.0484 | 0.0000 |
| Complexity | 15 | 100\% | 0.020 | 0.0000 | 0.0000 | 0.4531 | 0.0001 | 0.0000 | 0.0009 | 0.5459 | 0.0000 |
| CrossJurisdictional Activity | 14 | 100\% | 0.231 | 0.0000 | 0.0079 | 0.0000 | 0.0000 | 0.0000 | 0.0002 | 0.0000 | 0.9919 |

2013, $\Delta$ CoVar and the sum of in and out degrees seems able to reproduce the choice of the regulators. A similar conclusion, given a $p$-value above the $30 \%$, is improbable in 2014 even if the weight of eigenvector centrality is almost one. Analogous considerations derive from alternative analysis based on the categories. In fact, we get a coherent picture in 2013. In this year $\Delta$ CoVar has a weight over the $70 \%$ for substitutability. Also, it is largely participating in interconnectedness complexity and cross-jurisdictional activity. In + out degree is the only variable related to size and a relevant variable for interconnectedness. The role of PCA for complexity, cross-jurisdictional activity and interconnectedness is also clear. For 2014 instead, the alternative selections elect in degree as the most relevant variable. This variable is almost the only one relevant for substitutability, it is also important for complexity and size. Instead, eigenvector centrality is informative for complexity. In addition, we discover that $\Delta$ CoVar is relevant for size.

The graphs in Figure 2 show in the two years the empirical cumulative distribution functions for each risk measure in the two groups decided by the Regulator.

### 3.3. Robustness

In Table 4, we present some robustness check for the entire procedure. We considered 20 replications of the procedure with 200 hundred particles and 10000 random permutations in four settings:

1. Random start: random starting points of the swarm and random swarm update.
2. Equal weights: start with the swarm concentrated in the center of the unit hypercube and random swarm update. This corresponds to giving equal weights to all the measures.
3. Best Partial Test: start with the swarm concentrated on one edge of the unit hypercube with the $i$-th weight one for the measure with minimum $\lambda_{i}$ and zero for the others and

## Regulators selection, 2013



Regulators selection, 2014


Figure 2: Systemic Measures contribution to Selections
random swarm update. This is done to check if the procedure goes beyond the intuitive solution of choosing only the measure with the most significant partial test.
4. PSO const seed: the starting point of the swarm and the random number used for the updates are kept constant during the 20 replications.

We report the mean and the standard deviation of optimal $p$-values and the mean and standard deviation (in parentheses) of the weights. The set up are ranked in terms of ascending standard deviation of the $p$-values and ascending maximum standard deviation among the weights.
Differences among the four set up are not so relevant and we reach a precision on $p$-values below the fourth decimal digit and around the second digit for weights.
Finally, we remark that we also conducted robustness checks substituting the coefficient of variation to the Kolmogorov-Smirnov statistic in the computation of the test. The results with the coefficient of variation showed much more variability and we opted for the more robust Kolmogorov-Smirnov statistic.

Table 4: Performances based on Kolmogorov-Smirnov test statistic with $B=10000$ randomized permutations


## 4. Conclusions

We investigate the relationship between SIBs selections and equity-based systemic risk measures. The procedure proposed is completely nonparametric and requires only the hypothesis of exchangeability of variables. We optimize using particle swarms on the weights of a combination of permutation tests on the measures. The weights quantify how much the measure is informative about the selection. The optimal $p$-value measures how much the result is reliable. We consider regulator selections and selections based on macro categories. We found coherent results for the year 2013. All the selections consider $\Delta$ CoVar and the sum of in and out degrees as the most informative variables. For 2014 optimization on the regulator selection gave an unreliable result. But alternative selections points to in degree measure as the most informative. Although we consider our findings interesting we are obliged to point out the lack of regularity that they carry. Unfortunately, differences are already present in the numerosity of the selections associated with categories. In general, the requirement on single category seems much more stringent for 2014. Given the availability of only two years, it is difficult to understand where this variability comes from. The changes among the two years are quantitative and qualitative. Different categories are associated with the different systemic measure in different years. This feature hints maybe to a dynamical relationship among categories and systemic measures. It seems that the equity-based systemic risk measures are not able to track changes in the decision variables used by the regulator. In particular, for 2014, the measures are not informative about the interconnectedness category. In this light it would be important to include also the measures proposed by Diebold and Yılmaz (2014). Unfortunately, with 25 series, the curse of dimensionality does not allow to estimate the VAR needed for computing the measures. The use of some penalization technique as in recent Demirer, Diebold, Liu, and Yilmaz (2015) could solve the issue. It could also lead to better Granger Causality Network on which measures of Billio et al. (2012) are based.
Analogously refinement of MES (Brownlees and Engle (2010), Acharya, Engle, and Richardson (2012)) and CoVaR (Girardi and Ergün (2013)) based on Multivariate GARCH were proposed. Also, completely new equity-based measures are emerging from the literature. All those refinements and new measure should be in the future included in this kind of analysis. this preliminary analysis represents the first step of our research that would quest for better tracking measures.

## 5. Appendix

### 5.1. Descriptive Statistics

In Tables 5 and 6, we report the main Statistics of Stock Returns for the European Banks included in our analysis: Tukey five numbers, mean, standard deviation, Skewness and Kurtosis.

### 5.2. Systemic Risk Measures

In the following Table we report the names, the Bloomberg ticker and the associated systemic risk measures used in our analysis.

Table 5: Summary Statistics European Banks Stock Returns 2013

|  | Mean | Max | Min | S. D. | Skewness | Kurtosis | Median | Q5\% | Q95\% |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Banca Monte dei Paschi di Siena SpA (BMPS IM) | -0.001 | 0.137 | -0.099 | 0.033 | 0.584 | 5.339 | -0.002 | -0.051 | 0.061 |
| Barclays PLC (BARC LN) | 0.000 | 0.078 | -0.066 | 0.017 | 0.046 | 4.740 | 0.001 | -0.028 | 0.027 |
| Banco Bilbao Vizcaya Argentaria SA (BBVA SM) | 0.001 | 0.047 | -0.049 | 0.017 | -0.193 | 3.489 | 0.001 | -0.028 | 0.029 |
| Bankia (BKIA SM) | -0.005 | 1.066 | -0.922 | 0.130 | 0.111 | 35.991 | 0.000 | -0.145 | 0.083 |
| BNP Paribas SA (BNP FP) | 0.001 | 0.052 | -0.051 | 0.018 | -0.134 | 3.691 | 0.001 | -0.030 | 0.031 |
| Commerzbank AG (CBK GR) | 0.000 | 0.154 | -0.094 | 0.030 | 0.612 | 6.649 | -0.002 | -0.049 | 0.046 |
| Credit Agricole SA (ACA FP) | 0.002 | 0.067 | -0.060 | 0.021 | -0.009 | 4.119 | 0.002 | -0.032 | 0.035 |
| Danske Bank A/S (DANSKE DC) | 0.001 | 0.089 | -0.063 | 0.016 | 0.476 | 7.268 | 0.001 | -0.024 | 0.025 |
| Deutsche Bank AG (DBK GR) | 0.000 | 0.041 | -0.050 | 0.016 | -0.431 | 3.430 | 0.001 | -0.027 | 0.025 |
| DNB ASA (DNB NO) | 0.001 | 0.075 | -0.071 | 0.018 | 0.174 | 5.987 | 0.002 | -0.026 | 0.029 |
| Erste Group Bank AG (EBS AV) | 0.000 | 0.060 | -0.117 | 0.023 | -0.509 | 5.676 | 0.000 | -0.032 | 0.037 |
| Svenska Handelsbanken AB (SHBA SS) | 0.001 | 0.035 | -0.058 | 0.014 | -0.650 | 4.661 | 0.002 | -0.025 | 0.023 |
| HSBC Holdings PLC (HSBA LN) | 0.000 | 0.038 | -0.040 | 0.012 | 0.011 | 4.293 | 0.000 | -0.016 | 0.018 |
| ING Groep NV (INGA NA) | 0.001 | 0.054 | -0.066 | 0.020 | -0.207 | 3.493 | 0.002 | -0.036 | 0.034 |
| Intesa Sanpaolo SpA (ISP IM) | 0.001 | 0.062 | -0.095 | 0.023 | -0.437 | 4.476 | 0.001 | -0.033 | 0.035 |
| KBC Groep NV (KBC BB) | 0.002 | 0.069 | -0.072 | 0.022 | 0.036 | 3.804 | 0.002 | -0.029 | 0.037 |
| La Caixa Bank SA (CABK SM) | 0.001 | 0.111 | -0.105 | 0.023 | -0.043 | 7.160 | 0.001 | -0.033 | 0.036 |
| Lloyds Banking Group PLC (LLOY LN) | 0.002 | 0.079 | -0.039 | 0.017 | 0.539 | 4.453 | 0.000 | -0.024 | 0.030 |
| Nordea Bank AB (NDA SS) | 0.001 | 0.045 | -0.046 | 0.015 | -0.428 | 3.878 | 0.002 | -0.025 | 0.024 |
| Royal Bank of Scotland Group PLC (RBS LN) | 0.000 | 0.061 | -0.103 | 0.023 | -0.468 | 4.684 | 0.000 | -0.034 | 0.040 |
| Banco Santander SA (SAN SM) | 0.000 | 0.050 | -0.059 | 0.018 | -0.224 | 4.025 | 0.000 | -0.029 | 0.029 |
| Skandinaviska Enskilda Banken AB (SEBA SS) | 0.002 | 0.041 | -0.053 | 0.015 | -0.562 | 4.451 | 0.003 | -0.020 | 0.023 |
| Societe Generale SA (GLE FP) | 0.001 | 0.099 | -0.062 | 0.023 | 0.184 | 4.354 | 0.002 | -0.036 | 0.037 |
| Standard Chartered PLC (STAN LN) | -0.001 | 0.041 | -0.068 | 0.016 | -0.534 | 4.872 | 0.000 | -0.024 | 0.023 |
| Swedbank AB (SWEDA SS) | 0.001 | 0.095 | -0.082 | 0.017 | -0.220 | 9.011 | 0.002 | -0.026 | 0.024 |
| UniCredit S.p.A. (UCG IM) | 0.001 | 0.055 | -0.088 | 0.024 | -0.538 | 4.360 | 0.002 | -0.039 | 0.038 |

Table 6: Summary Statistics European Banks Stock Returns 2014

|  | Mean | Max | Min | S. D. | Skewness | Kurtosis | Median | Q5\% | Q95\% |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Banca Monte dei Paschi di Siena SpA (BMPS IM) | -0.004 | 0.346 | -0.242 | 0.053 | 0.538 | 15.134 | -0.002 | -0.077 | 0.059 |
| Barclays PLC (BARC LN) | 0.000 | 0.085 | -0.063 | 0.018 | 0.421 | 6.687 | 0.000 | -0.026 | 0.026 |
| Banco Bilbao Vizcaya Argentaria SA (BBVA SM) | 0.000 | 0.066 | -0.057 | 0.017 | 0.049 | 4.457 | -0.001 | -0.029 | 0.026 |
| Bankia (BKIA SM) | 0.000 | 0.080 | -0.062 | 0.021 | 0.241 | 3.639 | 0.000 | -0.033 | 0.035 |
| BNP Paribas SA (BNP FP) | -0.001 | 0.044 | -0.047 | 0.016 | 0.052 | 3.219 | 0.000 | -0.026 | 0.027 |
| Commerzbank AG (CBK GR) | 0.000 | 0.088 | -0.063 | 0.022 | 0.296 | 4.690 | 0.001 | -0.036 | 0.034 |
| Credit Agricole SA (ACA FP) | 0.001 | 0.073 | -0.060 | 0.020 | 0.040 | 4.109 | 0.000 | -0.029 | 0.031 |
| Danske Bank A/S (DANSKE DC) | 0.001 | 0.056 | -0.034 | 0.013 | 0.423 | 4.536 | 0.001 | -0.020 | 0.023 |
| Deutsche Bank AG (DBK GR) | -0.001 | 0.041 | -0.044 | 0.016 | -0.092 | 2.782 | -0.001 | -0.028 | 0.025 |
| DNB ASA (DNB NO) | 0.000 | 0.055 | -0.064 | 0.016 | -0.177 | 5.582 | 0.001 | -0.027 | 0.025 |
| Erste Group Bank AG (EBS AV) | -0.001 | 0.059 | -0.179 | 0.025 | -1.968 | 14.977 | 0.000 | -0.036 | 0.034 |
| Svenska Handelsbanken AB (SHBA SS) | 0.000 | 0.035 | -0.052 | 0.012 | -0.296 | 4.876 | 0.001 | -0.018 | 0.020 |
| HSBC Holdings PLC (HSBA LN) | 0.000 | 0.031 | -0.037 | 0.011 | -0.187 | 3.496 | 0.000 | -0.018 | 0.019 |
| ING Groep NV (INGA NA) | 0.000 | 0.056 | -0.058 | 0.018 | -0.121 | 3.623 | 0.000 | -0.032 | 0.029 |
| Intesa Sanpaolo SpA (ISP IM) | 0.001 | 0.068 | -0.064 | 0.022 | -0.023 | 3.155 | 0.002 | -0.036 | 0.041 |
| KBC Groep NV (KBC BB) | 0.001 | 0.060 | -0.060 | 0.019 | 0.045 | 3.805 | 0.000 | -0.031 | 0.034 |
| La Caixa Bank SA (CABK SM) | 0.001 | 0.094 | -0.058 | 0.020 | 0.270 | 4.643 | 0.001 | -0.033 | 0.032 |
| Lloyds Banking Group PLC (LLOY LN) | 0.000 | 0.055 | -0.050 | 0.015 | -0.303 | 4.132 | 0.001 | -0.028 | 0.021 |
| Nordea Bank AB (NDA SS) | 0.000 | 0.050 | -0.036 | 0.014 | 0.082 | 3.629 | 0.000 | -0.026 | 0.024 |
| Royal Bank of Scotland Group PLC (RBS LN) | 0.001 | 0.104 | -0.081 | 0.020 | 0.922 | 9.680 | 0.000 | -0.026 | 0.026 |
| Banco Santander SA (SAN SM) | 0.000 | 0.049 | -0.048 | 0.015 | -0.145 | 3.920 | 0.001 | -0.026 | 0.024 |
| Skandinaviska Enskilda Banken AB (SEBA SS) | 0.000 | 0.051 | -0.050 | 0.013 | -0.224 | 4.729 | 0.000 | -0.021 | 0.022 |
| Societe Generale SA (GLE FP) | -0.001 | 0.054 | -0.056 | 0.019 | 0.238 | 3.530 | -0.002 | -0.028 | 0.035 |
| Standard Chartered PLC (STAN LN) | -0.001 | 0.042 | -0.094 | 0.015 | -1.025 | 9.597 | -0.001 | -0.020 | 0.024 |
| Swedbank AB (SWEDA SS) | 0.000 | 0.046 | -0.068 | 0.013 | -0.465 | 6.987 | 0.001 | -0.019 | 0.019 |
| UniCredit S.p.A. (UCG IM) | 0.000 | 0.067 | -0.064 | 0.023 | 0.117 | 3.374 | -0.001 | -0.037 | 0.043 |

Table 7: Equity based systemic risk measures for European Banks based on stock returns in the year 2013. Transformations are applied in such a way that higher measure correspond to higher risk

|  | -MES | $-\Delta \mathrm{CoVaR}$ | in degree | out degree | in+out degree | 1/closeness | eigenvector centrality | PCA |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Banca Monte dei Paschi di Siena SpA (BMPS IM) | 0.0286 | 0.0057 | 1 | 1 | 2 | 0.429 | 0.381 | 0.040 |
| Barclays PLC (BARC LN) | 0.0268 | 0.0099 | 2 | 0 | 2 | 0.037 | 0.000 | 0.011 |
| Banco Bilbao Vizcaya Argentaria SA (BBVA SM) | 0.0248 | 0.0134 | 0 | 0 | 0 | 0.037 | 0.000 | 0.011 |
| Bankia (BKIA SM) | 0.0927 | 0.0016 | 3 | 1 | 4 | 1.000 | 0.127 | 0.628 |
| BNP Paribas SA (BNP FP) | 0.0358 | 0.0134 | 1 | 0 | 1 | 0.037 | 0.000 | 0.012 |
| Commerzbank AG (CBK GR) | 0.0302 | 0.0058 | 1 | 0 | 1 | 0.037 | 0.000 | 0.032 |
| Credit Agricole SA (ACA FP) | 0.0343 | 0.0106 | 1 | 1 | 2 | 1.000 | 0.000 | 0.017 |
| Danske Bank A/S (DANSKE DC) | 0.0175 | 0.0072 | 0 | 1 | 1 | 1.000 | 0.000 | 0.010 |
| Deutsche Bank AG (DBK GR) | 0.0217 | 0.0065 | 1 | 1 | 2 | 1.000 | 0.127 | 0.009 |
| DNB ASA (DNB NO) | 0.0291 | 0.0072 | 4 | 1 | 5 | 0.667 | 0.127 | 0.012 |
| Erste Group Bank AG (EBS AV) | 0.0405 | 0.0077 | 1 | 0 | 1 | 0.037 | 0.000 | 0.019 |
| Svenska Handelsbanken AB (SHBA SS) | 0.0234 | 0.0116 | 2 | 2 | 4 | 0.455 | 0.127 | 0.008 |
| HSBC Holdings PLC (HSBA LN) | 0.0183 | 0.0088 | 0 | 1 | 1 | 0.353 | 0.127 | 0.005 |
| ING Groep NV (INGA NA) | 0.0358 | 0.0087 | 0 | 1 | 1 | 0.429 | 0.381 | 0.015 |
| Intesa Sanpaolo SpA (ISP IM) | 0.0365 | 0.0075 | 0 | 3 | 3 | 0.464 | 0.508 | 0.019 |
| KBC Groep NV (KBC BB) | 0.0319 | 0.0063 | 0 | 2 | 2 | 0.571 | 0.127 | 0.018 |
| La Caixa Bank SA (CABK SM) | 0.0193 | 0.0058 | 1 | 1 | 2 | 0.667 | 0.127 | 0.019 |
| Lloyds Banking Group PLC (LLOY LN) | 0.0153 | 0.0070 | 1 | 2 |  | 0.556 | 0.127 | 0.010 |
| Nordea Bank AB (NDA SS) | 0.0227 | 0.0093 | 0 | 0 | 0 | 0.037 | 0.000 | 0.008 |
| Royal Bank of Scotland Group PLC (RBS LN) | 0.0279 | 0.0091 | 0 | 2 | 2 | 0.345 | 0.381 | 0.020 |
| Banco Santander SA (SAN SM) | 0.0274 | 0.0123 | 0 | 0 | 0 | 0.037 | 0.000 | 0.011 |
| Skandinaviska Enskilda Banken AB (SEBA SS) | 0.0267 | 0.0093 | 2 | 1 | 3 | 0.500 | 0.127 | 0.008 |
| Societe Generale SA (GLE FP) | 0.0385 | 0.0081 | 1 | 0 | 1 | 0.037 | 0.000 | 0.019 |
| Standard Chartered PLC (STAN LN) | 0.0174 | 0.0099 | 3 | 5 | 8 | 0.667 | 0.381 | 0.009 |
| Swedbank AB (SWEDA SS) | 0.0284 | 0.0116 | 2 | 0 | 2 | 0.037 | 0.000 | 0.011 |
| UniCredit S.p.A. (UCG IM) | 0.0382 | 0.0119 | 0 | 1 | 1 | 0.500 | 0.127 | 0.021 |

Table 8: Equity based systemic risk measures for European Banks based on stock returns in the year 2014 . Transformations are applied in such a way that higher measure correspond to higher risk

|  | -MES | - $\Delta \mathrm{CoVaR}$ | in degree | out degree | in + out degree | 1/closeness | eigenvector centrality | PCA |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Banca Monte dei Paschi di Siena SpA (BMPS IM) | 0.0308 | 0.0059 | 0 | 0 | 0 | 0.037 | 0.000 | 0.261 |
| Barclays PLC (BARC LN) | 0.0253 | 0.0096 | 1 | 0 | 1 | 0.037 | 0.000 | 0.030 |
| Banco Bilbao Vizcaya Argentaria SA (BBVA SM) | 0.0271 | 0.0132 | 1 | 9 | 10 | 0.600 | 0.672 | 0.026 |
| Bankia (BKIA SM) | 0.0255 | 0.0108 | 0 | 1 | 1 | 1.000 | 0.000 | 0.042 |
| BNP Paribas SA (BNP FP) | 0.0283 | 0.0133 | 3 | 3 | 6 | 0.545 | 0.000 | 0.024 |
| Commerzbank AG (CBK GR) | 0.0275 | 0.0071 | 3 | 0 | 3 | 0.037 | 0.000 | 0.046 |
| Credit Agricole SA (ACA FP) | 0.0267 | 0.0105 | 2 | 5 | 7 | 0.333 | 0.081 | 0.036 |
| Danske Bank A/S (DANSKE DC) | 0.0192 | 0.0099 | 11 | 0 | 11 | 0.037 | 0.000 | 0.016 |
| Deutsche Bank AG (DBK GR) | 0.0250 | 0.0127 | 1 | 1 | 2 | 1.000 | 0.000 | 0.024 |
| DNB ASA (DNB NO) | 0.0243 | 0.0095 | 3 | 2 | 5 | 0.750 | 0.000 | 0.025 |
| Erste Group Bank AG (EBS AV) | 0.0308 | 0.0101 | 1 | 7 | 8 | 0.512 | 0.416 | 0.057 |
| Svenska Handelsbanken AB (SHBA SS) | 0.0154 | 0.0114 | 3 | 0 | 3 | 0.037 | 0.000 | 0.013 |
| HSBC Holdings PLC (HSBA LN) | 0.0182 | 0.0105 | 7 | 1 | 8 | 0.500 | 0.000 | 0.012 |
| ING Groep NV (INGA NA) | 0.0322 | 0.0131 | 1 | 1 | 2 | 1.000 | 0.000 | 0.031 |
| Intesa Sanpaolo SpA (ISP IM) | 0.0297 | 0.0136 | 3 | 9 | 12 | 0.553 | 0.503 | 0.047 |
| KBC Groep NV (KBC BB) | 0.0303 | 0.0082 | 1 | 3 | 4 | 0.362 | 0.156 | 0.035 |
| La Caixa Bank SA (CABK SM) | 0.0265 | 0.0128 | 0 | 1 | 1 | 1.000 | 0.000 | 0.037 |
| Lloyds Banking Group PLC (LLOY LN) | 0.0238 | 0.0138 | 2 | 1 | 3 | 1.000 | 0.000 | 0.021 |
| Nordea Bank AB (NDA SS) | 0.0198 | 0.0158 | 3 | 0 | 3 | 0.037 | 0.000 | 0.020 |
| Royal Bank of Scotland Group PLC (RBS LN) | 0.0239 | 0.0072 | 2 | 0 | 2 | 0.037 | 0.000 | 0.036 |
| Banco Santander SA (SAN SM) | 0.0259 | 0.0133 | 4 | 8 | 12 | 0.477 | 0.302 | 0.021 |
| Skandinaviska Enskilda Banken AB (SEBA SS) | 0.0217 | 0.0115 | 1 | 0 | 1 | 0.037 | 0.000 | 0.016 |
| Societe Generale SA (GLE FP) | 0.0349 | 0.0099 | 1 | 1 | 2 | 0.385 | 0.000 | 0.035 |
| Standard Chartered PLC (STAN LN) | 0.0184 | 0.0108 | 0 | 1 | 1 | 0.500 | 0.000 | 0.020 |
| Swedbank AB (SWEDA SS) | 0.0162 | 0.0122 | 1 | 2 | 3 | 0.750 | 0.000 | 0.015 |
| UniCredit S.p.A. (UCG IM) | 0.0349 | 0.0126 | 3 | 2 | 5 | 0.500 | 0.000 | 0.052 |

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[^0]:    ${ }^{1}$ This research used the SCSCF multiprocessor cluster system at University Ca' Foscari of Venice.

